Survival Function Estimation with Recurrent Events: Case of Retinal Neural Firing

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## Introduction

Outline:

Introduction to Survival Analysis



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- Introduction to Survival Analysis
- O Survival Analysis with a Single Event



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Introduction to Survival Analysis

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- O Survival Analysis with a Single Event
- O How to Estimate the Survival Function of Recurrent Events



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- Introduction to Survival Analysis
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- Introduction to Survival Analysis
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- Analysis of our Project Data
- Future Work

Conclusion

### What is Survival Analysis?

• Survival analysis is a field of statistics that analyzes and models time-to-event data.



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- This area of statistics is widely used in medical research, economics, and reliability.
- Examples:
  - Time from cancer remission to relapse.
  - Duration between administering a treatment and recovery.

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#### Special Features of Survival Data

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  - Individuals whose survival times can not be analyzed because they have been lost to follow-up.
  - Occurance of an event due to a cause other than one of interest.



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- Let r be the number of event times amongst the individuals, so that r ≤ n, where r does not include censored times and duplicate times are considered as one time.
- Now let  $t_{(j)}$ , for j = 1, 2, 3, ..., r, be the r ordered event times such that  $t_{(1)} < t_{(2)} < ... t_{(r)}$ .

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- Let  $n_j$  be the number of individuals at risk just before  $t_{(j)}$ .
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- The survival function, S(t), is the probability that the time of event is later than some specified time t.

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• A common way of estimating the survival function is the Kaplan-Meier estimator. (Kaplan, E.L & Meier P., 1958)

#### The Kaplan-Meier Estimator

#### • Kaplan-Meier (Product Limit) Estimate:

$$\widehat{S}(t) = \prod_{j=1}^{r} \left\{ 1 - \frac{d_j}{n_j} \right\}$$
  
SE  $\widehat{S}(t) \approx \widehat{S}(t) \left\{ \sum_{j=1}^{r} \frac{d_j}{n_j(n_j - d_j)} \right\}^{\frac{1}{2}}$ 



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•  $(1 - \alpha) * 100\%$  Confidence Intervals of  $\widehat{S}(t)$ :  $\widehat{S}(t) \pm z_{\alpha/2} * \text{SE } \widehat{S}(t)$ 

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## Illustrative Example of Kaplan-Meier

#### Example

Suppose there is a sample of 12 hemophiliacs under the age of 40 with HIV seroconversion.

Instead of using time intervals, exact times at which failures occured are used.

These times, in months, are listed below:

 Patient
 1
 2
 3
 4
 5
 6
 7
 8
 9
 10
 11
 12

 Months
 2
 3\*
 6
 6
 7
 10\*
 15
 16
 27
 30
 32

NOTE: \* denotes censored values.





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The purpose of our project is to estimate the survival function of recurrent event data.



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- When multiple events occur within the same subject, they are known as **recurrent events**.
- Examples include migraines, seizures, heart attacks, strokes etc.



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#### Time to Event Plot



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- Let  $T_{ij}$  be the time from the  $(j-1)^{th}$  to the  $j^{th}$  event for subject *i*.

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- Let  $T_{ij}$  be the time from the  $(j-1)^{th}$  to the  $j^{th}$  event for subject *i*.
- Let the censoring time,  $C_i$ , be the time between the initial event and the end of follow-up for subject *i*.

#### • Let *m<sub>i</sub>* denote the index satisfying:

$$\sum_{j=1}^{m_i-1} T_{ij} \leq C_i$$

and

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• Let  $m_i$  be the number of recurrent events for subject *i*.  $m_i^*$  is the number of uncensored recurrent events for subject *i*.

$$m_i^* = \begin{cases} 1 & \text{if } m_i = 1 \\ m_i - 1 & \text{if } m_i \ge 2 \end{cases}$$

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• Let  $y_{ij}$  be our observed recurrence times defined by:

$$y_{ij} = \left\{ egin{array}{cc} t_{ij} & ext{if} & j=1,\ldots,m_i-1 \ t^+_{i,m_i} & ext{if} & j=m_i \end{array} 
ight.$$

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- Terminology
  - Let  $R^*(t)$  be the total mass of the risk set at time t be calculated as:

$$R^*(t) = \sum_{i=1}^n \left[ \frac{a_i}{m_i^*} \sum_{j=1}^{m_i^*} I(y_{ij} \ge t) \right]$$

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 Let a<sub>i</sub> be defined as a positive-valued function of the censored value subject to the constraint E(a<sub>i</sub><sup>2</sup>) < ∞.</li>

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- The indicator function,  $I(y_{ij} \ge t)$ , is a binary operator with values 0 if  $y_{ij} < t$  and 1 if  $y_{ij} \ge t$ .

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- The indicator function,  $I(y_{ij} \ge t)$ , is a binary operator with values 0 if  $y_{ij} < t$  and 1 if  $y_{ij} \ge t$ .
- $R^*(t)$  is the summation of the weighted average of the total number of observed uncensored recurrent times for a subject that are greater than or equal to t.

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• Let the mass evaluated at time t be:

$$d^{*}(t) = \sum_{i=1}^{n} \left[ \frac{a_{i}I(m_{i} \ge 2)}{m_{i}^{*}} \sum_{j=1}^{m_{i}^{*}} I(y_{ij} = t) \right]$$



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•  $d^*(t)$  is the summation of the weighted average of the total number of observed uncensored recurrent times for a subject that are equal to t.

## Wang and Chang Product-Limit Estimation (1999)

 Letting y<sub>1</sub><sup>\*</sup>, y<sub>2</sub><sup>\*</sup>, ..., y<sub>K</sub><sup>\*</sup> be the ordered, and distinct uncensored times, Wang and Chang created a Kaplan-Meier type estimator:

$$\widehat{S}_n(t) = \prod_{y_i^* \leq t} \left\{ 1 - \frac{d^*(y_i^*)}{R^*(y_i^*)} \right\}$$



# Wang and Chang Product-Limit Estimation (1999)

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$$\widehat{S}_n(t) = \prod_{y_i^* \leq t} \left\{ 1 - rac{d^*(y_i^*)}{R^*(y_i^*)} \right\}$$

• This estimator also sets *a<sub>i</sub>* equal to 1, giving every estimate equal weight.

## Product-Limit Estimate



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• Using the R package **survrec**, we found the median survival time, *t*<sub>.5</sub>, to be 16.6 ms.



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- Using the R package **survrec**, we found the median survival time, *t*<sub>.5</sub>, to be 16.6 ms.
- The 95% confidence interval around the median survival time is (0.4693057, 32.7306943).
  - The standard error was found through Greenwood's formula. (Greenwood M., 1926)

Single Event Analysis

Recurrent Event Analysis

Data Analysis

Conclusion

#### WC Plot for 2 Inter-Event Times:

Median Survival Time: 138 ms



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Single Event Analysis

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#### WC Plot for 3 Inter-Event Times:

Median Survival Time: 35 ms



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#### WC Plot for 4 Inter-Event Times:

Median Survival Time: 21.5 ms



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Conclusion

### Future Work

#### Latency



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## Future Work

- Latency
- Correlation



## Future Work

- Latency
- Correlation
- More Subjects



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- University of Iowa, Department of Ophthamology and Neurology
- National Heart Lung and Blood Institute (NHLBI)
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### References

- Greenwood, M. 1946. The Statistical Study of Infectious Diseases. *Journal of the Royal Statistical Society* 109(2):85-110
- Kaplan, E.L., Meier, P. 1958. Nonparametric estimation from incomplete observations. *Journal of the American Statistical Association* 53:457–481
- - Wang, M.C., Chang, S.H. 1999. Nonparametric Estimation of a Recurrent Survival Function. *Journal of the American Statistical Association* 94:146–153.